


Shared genetic components between metabolic syndrome and schizophrenia: Genetic correlation using multipopulation data sets

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Aim: The genetic relationship between schizophrenia (SCZ) and other nonpsychiatric disorders remains largely unknown. We examined the shared genetic components between these disorders based on multipopulation data sets.

Methods: We used two data sets for East Asian (EAS) and European (EUR) samples. SCZ data was based on the Psychiatric Genomics Consortium Asia with our own genome-wide association study for EAS and Psychiatric Genomics Consortium for EUR. Nonpsychiatric data (20 binary traits [mainly nonpsychiatric complex disorders] and 34 quantitative traits [mainly laboratory examinations and physical characteristics]) were obtained from Biobank Japan and UK Biobank for EAS and EUR samples, respectively. To evaluate genetic correlation, linkage disequilibrium score regression analysis was utilized with further meta-analysis for each result from EAS and EUR samples to obtain robust evidence. Subsequent mendelian randomization analysis was also included to examine the causal effect.

Results: A significant genetic correlation between SCZ and several metabolic syndrome (MetS) traits was detected in the combined samples (meta-analysis between EAS and EUR data) (body mass index [$r_g = -0.10$, $q\text{-value} = 1.0 \times 10^{-9}$], high-density-lipoprotein cholesterol [$r_g = 0.072$, $q\text{-value} = 2.9 \times 10^{-3}$],

blood sugar [$r_g = -0.068$, $q\text{-value} = 1.4 \times 10^{-2}$], triglycerides [$r_g = -0.052$, $q\text{-value} = 2.4 \times 10^{-2}$], systolic blood pressure [$r_g = -0.054$, $q\text{-value} = 3.5 \times 10^{-2}$], and C-reactive protein [$r_g = -0.076$, $q\text{-value} = 7.8 \times 10^{-5}$]. However, no causal relationship on SCZ susceptibility was detected for these traits based on the mendelian randomization analysis.

Conclusion: Our results indicate shared genetic components between SCZ and MetS traits and C-reactive protein. Specifically, we found it interesting that the correlation between MetS traits and SCZ was the opposite of that expected from clinical studies: this genetic study suggests that SCZ susceptibility was associated with reduced MetS. This implied that MetS in patients with SCZ was not associated with genetic components but with environmental factors, including antipsychotics, lifestyle changes, poor diet, lack of exercise, and living conditions.

Keywords: blood pressure, body mass index, cholesterol, C-reactive protein, triglycerides.

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Schizophrenia (SCZ) is a common disorder with a complex pathophysiology. Although numerous studies have aimed to elucidate the biological causes,^{1–7} little evidence has been available so far. One reason is the ambiguous diagnostic boundaries between SCZ and other psychiatric disorders⁸ given that current diagnostic systems have mainly been based on descriptive criteria proposed by psychopathology. Although such criteria, e.g., the *DSM* or *ICD* systems, provide a consensus to some extent, classical epidemiological studies have estimated the shared genetic liability among psychiatric disorders.^{9–12}

To support this, recent advances in genomic analysis have provided important insight that would help address whether a genetic cause is shared between SCZ and other psychiatric disorders or whether it is unique to each disorder, based on data from association results of whole-genome single nucleotide polymorphisms (SNPs). Specifically, the cross-disorder group of the Psychiatric Genomics Consortium (PGC) reported a significant genetic overlap among them based on genetic correlation analysis and polygenic risk score analysis.^{13–15} Interestingly, the genetic

component of SCZ is close to that of bipolar disorder, of which contribution is larger than that of the relationship between bipolar disorder and major depressive disorder.^{6,8,14,15}

To determine the pathophysiology of SCZ and promote genome drug discovery, as well as prioritize the possible causality, understanding the genetic correlation between psychiatric disorders and/or other nonpsychiatric traits (disorders and quantitative traits) is crucial.¹⁶ Specifically, nonpsychiatric traits can help uncover relationships that have yet to be reported or have been overlooked. For instance, several studies have reported that the genetic component for lower body mass index (BMI) correlated with that for susceptibility to SCZ.^{16–18} Moreover, genetic variants associated with high C-reactive protein (CRP) level (an instrumental variable) as a modifiable exposure to SCZ showed significant causality through mendelian randomization (MR) analysis, which is usually applied with genetic correlation analysis,^{19–21} although some studies did not show a causal relationship.^{22,23}

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Here, we examined the genetic relationship between SCZ and other nonpsychiatric disorders based on linkage disequilibrium score regression (LDSC) analysis, which is suitable for examining disorders with polygenic architecture.¹⁶ Furthermore, we included subsequent analysis for the causal relationship between significant instrumental variants in nonpsychiatric traits and SCZ based on MR analysis.²⁴ The current study highlights the analyses using data from multiple populations. Although previous studies have mainly analyzed data based on European (EUR) populations, multipopulation analysis can provide more robust results. Therefore, we used data based on East Asian (EAS) populations in addition to that of EUR populations and finally assessed LDSC results using meta-analysis to enhance statistical power.

Methods

Genome-wide association study results in EAS and EUR samples

For EAS samples (a total of 24,504 patients with SCZ and 42,770 controls), we used SCZ results based on the meta-analysis from our previous study (Japanese SCZ: 1,726 patients with SCZ and 7,408 controls)⁶ and the PGC Asia (PGC EAS SCZ: 22,778 patients with SCZ and 35,362 controls),²⁵ as well as nonpsychiatric data from the JENGER database (40 binary traits and 63 quantitative traits: Supplementary Table 1), which is provided by Biobank Japan (<http://jenger.riken.jp/result>) (minimum of 12,302 and maximum of 212,453 samples). The nonpsychiatric traits were common diseases (e.g., type 2 diabetes), quantitative data of laboratory examinations (e.g., blood sugar [BS]), and physical characteristics (e.g., BMI). It should be noted that there were no overlapping samples between GWASs for SCZ and nonpsychiatric traits.

To enhance the statistical power during the primary meta-analysis of SCZ results for EAS, the “metafor” package in R (<https://www.metafor-project.org/doku.php>) was used. We applied results from the fixed-effects model with $P_{\text{heterogeneity}} \geq 0.05$ and the random-effects model $P_{\text{heterogeneity}} < 0.05$.

For EUR samples, we used PGC results for SCZ (33,640 SCZ and 43,456 controls)²⁶ and UK Biobank data (<http://www.nealelab.is/uk-biobank>: a total of 4,236 phenotypes are listed) for nonpsychiatric traits (minimum of 51,453 and maximum of 361,194 samples). Among these nonpsychiatric traits, we used overlapped ones with the phenotypes listed in JENGER (20 binary traits and 34 quantitative traits).

LDSC analysis

All genetic correlation analyses were conducted using LDSC¹⁶ software (<https://github.com/bulik/ldsc>). This software estimates the genetic correlation between two phenotypes based on the summary statistics. The genetic correlations between SCZ and nonpsychiatric traits were calculated in each population (EAS and EUR). However, specifically for EUR data, we only assessed the genetic correlation in identical pairs analyzed from EAS results, thereby excluding traits listed only in the UK Biobank, mentioned above.

For quality control, we used SNPs: (1) involved in the 1000 Genomes Project and HapMap Projects, (2) with a minor allele frequency $\geq 1\%$, and (3) INFO (imputation quality measures) ≥ 0.9 .

Subsequent meta-analysis was conducted by merging the “overlapped” traits (20 binary traits and 34 quantitative trait loci [QTL] traits) between results from the EAS and EUR samples. The “metafor” package in R was used based on the fixed-effects model, because we assumed little differences of the genetic correlation between populations. Multiple comparisons were corrected using the false discovery rate (Benjamini–Hochberg method) to calculate q -values (q -values < 0.05 were set as the significant threshold) in each trait (binary and QTL traits in the meta-analysis).

MR analysis

For MR analysis, we only targeted traits that showed significant correlation during LDSC analysis. These nonpsychiatric traits were used as

exposures, whereas SCZ was used as the outcome. The “exposure” SNPs, which surpassed the P -value threshold of $P < 5 \times 10^{-8}$ in the results of the nonpsychiatric traits, were clamped to remove SNPs in linkage disequilibrium ($r^2 < 0.001$ within a 10-Mb window) and selected as instrumental variables. Our primary MR analysis method was inverse variance weighted (IVW) regression, particularly when P -values of the Egger intercept exceeded 0.05. We also conducted: (1) weighted median analysis, (2) MR-Egger regression analysis, and (3) MR pleiotropy residual sum and outlier (MR-PRESSO) analysis as sensitivity analyses. During MR-PRESSO, we checked the horizontal pleiotropy using the robust method and performed the global test to examine for outlier removal.

The aforementioned MR analyses were conducted using the “TwoSample MR”²⁷ package in R (<https://mrcieu.github.io/TwoSampleMR/>), with the default setting. The multiple comparisons were again corrected by false discovery rate and the significance level was set at q -values < 0.05 in each population (EAS and EUR).

Ethical statement

For the SCZ GWAS in the Japanese population,⁶ written informed consent was obtained from all patients following a thorough explanation of the study. The study was approved by the ethics committees of Fujita Health University and other participating universities, which conform to the provisions of the Declaration of Helsinki.

Results

Initially, LDSC analysis was conducted between SCZ and nonpsychiatric traits in EAS samples (Tables 1 and Fig. 1; the full results can be seen in Table S1 and Figs. S1 and S2). Ten traits showed marginal significant correlation ($P < 0.05$) in this population (Table S1). Next, we conducted LDSC analysis for EUR samples and detected 8 of 54 “overlapped” traits (20 binary traits and 34 QTL traits) with $P < 0.05$ (Table S1). The subsequent meta-analysis for the 54 “overlapped” traits in each population revealed no significant genetic correlation between SCZ and binary traits (q -value < 0.05); however, we detected six QTL traits that showed significant genetic correlation with SCZ (Table 1 and Fig. 1): BMI ($r_g = -0.101$, q -value = 1.01×10^{-9}), CRP ($r_g = -0.0756$, q -value = 7.80×10^{-5}), high-density-lipoprotein cholesterol (HDL-C: $r_g = 0.0722$, q -value = 2.90×10^{-3}), BS ($r_g = -0.0679$, q -value = 1.44×10^{-2}), triglycerides (TGs: $r_g = -0.0524$, q -value = 2.44×10^{-2}), and systolic blood pressure (SBP: $r_g = -0.0543$, q -value = 3.45×10^{-2}). The direction of the genetic correlation for the six mentioned traits was identical between EAS and EUR results, suggesting consistent evidence between populations.

Further analysis on these “significant” traits was conducted through MR to examine causality. No significant causal relationship was obtained after correcting the multiple comparisons (Table 2). However, it is of note that we detected nonsignificant trend of the causal effects between CRP ($\beta = -0.0923$, q -value = 0.0514 in IVW, Egger intercept = 0.389) and SCZ, which is derived from EUR samples (Table 2). Whereas, the results with the significant P -value of Egger intercept ($P < 0.05$, which means possible pleiotropy, thus prioritizing other “non-IVW” methods) were not always consistent.

Discussion

In this genetic correlation analysis using multipopulation samples, we detected a significant correlation between several metabolic syndrome (MetS)-related traits/CRP and SCZ. Most of these have been previously reported mainly in EUR samples,^{16–18} suggesting that our results replicated those presented in previous studies, albeit in a larger sample size with multipopulation samples. Nevertheless, we did not detect the significant causal relationship between these nonpsychiatric traits and SCZ after correcting the multiple comparisons.

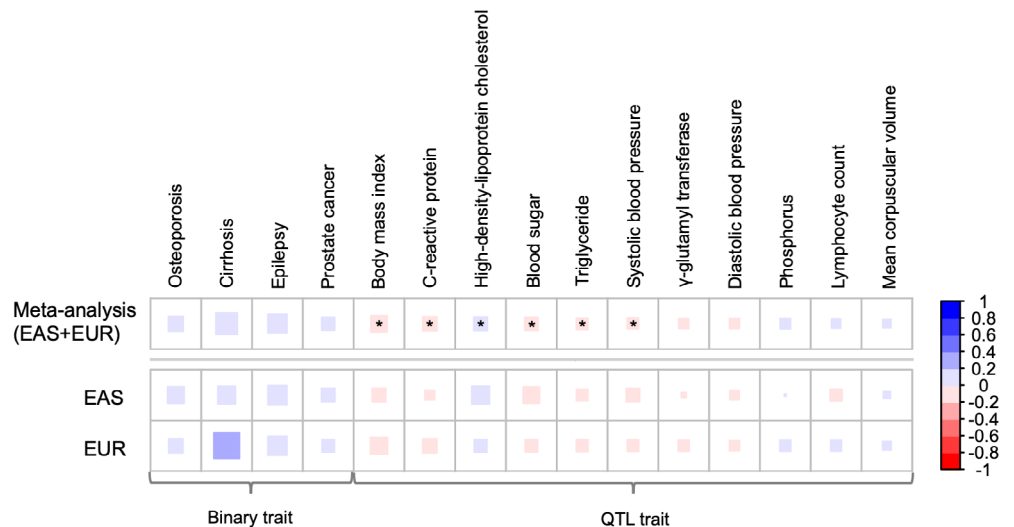
Table 1. Results of genetic correlation analysis with $P < 0.1$

Trait	Type	EAS			EUR			Meta-analysis (EAS + EUR)			
		r_g	SE	P -value	r_g	SE	P -value	r_g	SE	P -value	q-value
Osteoporosis	Binary	0.107	0.0784	0.173	0.0789	0.0439	0.0725	0.0856	0.0383	0.0254	0.377
Cirrhosis	Binary	0.121	0.121	0.320	0.254	0.141	0.0719	0.178	0.0737	0.0538	0.377
Epilepsy	Binary	0.140	0.126	0.268	0.141	0.0907	0.121	0.141	0.0920	0.0565	0.377
Prostate cancer	Binary	0.0703	0.0510	0.168	0.0594	0.0614	0.333	0.0659	0.0392	0.0933	0.467
BMI	QTL	-0.0713	0.0294	0.0155	-0.112	0.0178	2.80E-10	-0.101	0.0152	2.97E-11	1.01E-09
CRP	QTL	-0.0363	0.0936	0.698	-0.0772	0.0186	3.18E-05	-0.0756	0.0182	3.40E-05	7.80E-05
HDL-C	QTL	0.122	0.0467	9.00E-03	0.0614	0.0218	4.90E-03	0.0722	0.0198	2.55E-04	2.90E-03
BS	QTL	-0.103	0.0475	0.0304	-0.0588	0.0243	0.0154	-0.0679	0.0216	1.70E-03	0.0144
TGs	QTL	-0.0510	0.0344	0.139	-0.0529	0.0211	0.0121	-0.0524	0.0180	3.60E-03	0.0244
SBP	QTL	-0.0694	0.0415	0.0943	-0.0498	0.0225	0.0267	-0.0543	0.0198	6.10E-03	0.0345
GGT	QTL	-9.10E-03	0.0343	0.791	-0.0465	0.0203	0.0219	-0.0368	0.0175	0.0352	0.171
DBP	QTL	-0.0345	0.0428	0.420	-0.0385	0.0222	0.0834	-0.0377	0.0197	0.0561	0.221
Phosphorus	QTL	1.40E-03	0.0758	0.985	0.0460	0.0233	0.0483	0.0422	0.0223	0.0584	0.221
Lymphocyte count	QTL	-0.0545	0.0494	0.270	0.0449	0.0193	0.0198	0.0317	0.0180	0.0775	0.264
MCV	QTL	0.0190	0.0343	0.580	0.0279	0.0174	0.109	0.0261	0.0155	0.0928	0.287

Bold numbers represent a significant correlation after applying false discovery rate correction.

BMI, body mass index; BS, blood sugar; CRP, C-reactive protein; DBP, diastolic blood pressure; EAS, East Asian; EUR, European; GGT, γ -glutamyl transferase; HDL-C, high-density lipoprotein cholesterol; MCV, mean corpuscular volume; QTL, quantitative trait loci; r_g , genetic correlation; SBP, systolic blood pressure; SE, standard error; TGs, triglycerides; QTL, quantitative trait loci.

Fig. 1 Heatmap of genetic correlation analysis with P -values < 0.1 (for overall samples). The colored box indicates the genetic correlation matrix (r_g). Asterisk indicates q -values < 0.05 . EAS, East Asian; EUR, European.



MetS-related traits and SCZ

The current study found a significant genetic correlation between MetS-related traits, such as BMI, HDL-C, BS, TGs, SBP, and SCZ. MetS is diagnosed based on various factors, namely large abdominal circumference, high TGs, low HDL-C, high blood pressure, and high fasting blood glucose.²⁸ Most of our results implied shared genetic components between “non-MetS”-related traits and SCZ; we found a negative r_g between SCZ and BMI, TGs, SBP, and BS, and a positive r_g between SCZ and HDL-C, all of which showed the identical direction of the effect for EAS and EUR samples, suggesting robust results.

Nevertheless, in clinical practice, patients with SCZ generally demonstrate a higher prevalence of dyslipidemia, obesity, hypertension, and hyperglycemia than the general population and often have an increased risk of premature death caused by cardiovascular

disease.^{29,30} Some epidemiological studies targeting first episodes among drug-naïve patients reported that such individuals had already exhibited high BMI or dyslipidemia.^{31–33} This may imply that second-generation antipsychotics, which is a major risk factor for obesity, are “independent” of MetS-related traits in the early stages of SCZ.

On the contrary, longitudinal studies surveying BMI, one of the MetS-related traits, during the adolescent period (e.g., aged 18 years) showed that patients with SCZ had a lower BMI than patients without SCZ.^{34,35} This evidence is consistent with our current results and previous reports on genetic correlation, which supports the classical “somatotype” theory, suggesting that obesity in SCZ is an acquired/extrinsic effect from drug therapy (antipsychotics), lifestyle changes, poor diet, lack of exercise, and living conditions.¹⁸

Table 2. Results of mendelian randomization analysis

Trait	Method	No. of SNPs	EAS				No. of SNPs	EUR			
			beta	SE	P-value	q-value		beta	SE	P-value	q-value
BMI	Inverse variance weighted	62	-0.211	0.113	0.0617	0.434	298	-0.165	0.0654	0.0116	0.0514
	Weighted median	62	-0.135	0.115	0.241	0.530	298	-0.125	0.0727	0.0860	0.229
	MR Egger	62	0.178	0.364	0.625	0.772	298	0.280	0.191	0.142	0.341
	MR-PRESSO	60	-0.125	0.0949	0.192	0.530	278	-0.140	0.0561	0.0134	0.0514
	Egger intercept				0.264					0.0135	
CRP	Inverse variance weighted	6	-0.0530	0.141	0.706	0.795	183	-0.0923	0.0378	0.0146	0.0514
	Weighted median	6	-0.0920	0.143	0.520	0.772	183	-0.109	0.0422	0.0100	0.0514
	MR Egger	6	0.197	0.342	0.596	0.772	183	-0.0602	0.0530	0.258	0.442
	MR-PRESSO	6	NA	NA	NA	NA	176	-0.101	0.0332	2.79E-03	0.0514
	Egger intercept				0.465					0.389	
HDL-C	Inverse variance weighted	50	0.0756	0.0364	0.0379	0.431	247	0.0412	0.0387	0.287	0.459
	Weighted median	50	0.0567	0.0543	0.296	0.530	247	-0.0184	0.0439	0.674	0.887
	MR Egger	50	-6.90E-03	0.0656	0.917	0.917	247	-0.0786	0.0559	0.161	0.351
	MR-PRESSO	50	NA	NA	NA	NA	235	0.0406	0.0323	0.209	0.386
	Egger intercept				0.139				2.00E-03	3.68E-03	
BS	Inverse variance weighted	15	-0.117	0.105	0.265	0.530	93	-0.135	0.0726	0.0629	0.189
	Weighted median	15	-0.145	0.141	0.303	0.530	93	-0.0485	0.0694	0.485	0.728
	MR Egger	15	-0.124	0.529	0.818	0.859	93	-0.0421	0.118	0.723	0.887
	MR-PRESSO	15	NA	NA	NA	NA	86	-0.126	0.0503	0.0144	0.0514
	Egger intercept				0.989					0.321	
TGs	Inverse variance weighted	36	-0.0469	0.0625	0.452	0.730	204	-0.0143	0.0428	0.739	0.887
	Weighted median	36	-0.0960	0.0671	0.152	0.530	204	-0.0208	0.0444	0.640	0.887
	MR Egger	36	-0.207	0.0975	0.0409	0.431	204	0.0825	0.0616	0.182	0.364
	MR-PRESSO	34	-0.0626	0.0564	0.275	0.530	192	6.83E-03	0.0340	0.841	0.910
	Egger intercept				0.0452					0.0316	
SBP	Inverse variance weighted	22	0.207	0.164	0.206	0.530	158	-0.0134	0.0831	0.872	0.910
	Weighted median	22	0.0794	0.161	0.623	0.772	158	0.0212	0.0846	0.802	0.910
	MR Egger	22	0.987	0.574	0.101	0.530	158	0.702	0.268	9.56E-03	0.0514
	MR-PRESSO	21	0.0416	0.114	0.719	0.795	153	-5.01E-03	0.0687	0.942	0.942
	Egger intercept				0.172					5.65E-03	

Bold numbers represent results for prioritized analysis based on the significance of Egger intercept (if $P > 0.05$, we prioritized the “inverse variance-weighted” method, whereas if $P < 0.05$, we prioritized other methods).

†Mendelian randomization (MR) pleiotropy residual sum and outlier (MR-PRESSO) beta effects were calculated after removing the outliers, which were pleiotropic variants. If there was no outlier, we described “NA” in the column. The number of single nucleotide polymorphism (SNPs) in MR-PRESSO was that after removing the outliers.

BMI, body mass index; BS, blood sugar; CRP, C-reactive protein; EAS, East Asian; EUR, European; HDL-C, high-density lipoprotein cholesterol; SBP, systolic blood pressure; SE, standard error; TGs, triglycerides.

It is important to determine whether these MetS-related traits in SCZ are derived from endogenous or extrinsic effects.^{19,36,37} Our findings basically support the notion that extrinsic effects are the major cause for not only BMI but also other MetS-related traits in SCZ.²⁸ In clinical settings, this speculation is encouraging for psychiatrists and patients because it implies that MetS-related traits can be manageable. However, it is clear that MetS-related traits in SCZ are caused by complex mechanisms and require further studies.

Causal effect of instrumental variables from nonpsychiatric traits on SCZ

Our MR analysis revealed no significant causal relationship between nonpsychiatric traits (MetS-related traits and CRP) and SCZ after correcting the multiple comparisons (q -value > 0.05). This indicates that, at least, there is no “strong” causal effect based on these traits.

However, despite a nonsignificant trend, we stressed that one of the best P -values (q -value) was obtained in the instrumental variables

from CRP of the EUR population ($q\text{-value}_{IVW} = 0.0514$, Egger intercept = 0.389). This is not surprising because several studies have suggested that CRP is a potential biomarker for diagnosis, treatment response, and prevention of SCZ.³⁸ Specifically, previous clinical studies, including those with a cross-sectional design, have shown that patients with SCZ had higher CRP values compared with the control group.^{39–45} However, such a study design contains bias, which could be associated with the “status” of SCZ.⁴⁶

In contrast, recent evidence from MR analyses, which can theoretically reduce the bias caused by environmental factors, suggest that high CRP values had a protective effect against SCZ.^{20,21} In the current study, a negative correlation was observed between CRP and SCZ in the LDSC analysis, while a negative beta was observed during MR analysis (exposure: CRP, outcome: SCZ [note: it was not “significant”]), although such observations were only observed in the EUR sample. Therefore, our results are at least in line with those presented in previous studies, which showed that higher CRP level had a protective effect against SCZ in EUR populations.^{20,21} This is partially supported by evidence in EAS, which showed the same direction of causal effect, albeit not significantly. One possible reason for the non-significant trend of CRP on SCZ (in combined samples, EUR, and/or EAS samples) was that the sample size of EUR, as well as EAS, samples was not sufficient. We presume that the increased sample size may allow us to obtain “significant” causal effects; thus, further studies using multipopulation samples will be essential to obtain concrete results.

Limitations

There are several limitations to interpreting our current results. First, the sample size of GWASs, which were used in the LDSC and MR analyses, might not be sufficient for both the SCZ and nonpsychiatric traits. This means that insufficient statistical power overlooked the significant correlation and/or causal relationships. Specifically, MR analysis for EAS samples was typical because the numbers of instrumental variables were much lower than those for EUR samples. Therefore, further GWASs are required to deny the type II error. Second, in the LDSC and MR analyses for the EAS population, GWAS data for SCZ were mainly from Han Chinese and Japanese, whereas nonpsychiatric data were from Japanese. This might influence the results as a result of different ancestries (albeit not large compared with the difference between EAS and EUR). Therefore, we conducted further LDSC analyses to confirm whether our analysis in possible “stratified” EAS affects the result. We calculated the genetic correlation between nonpsychiatric traits (Biobank Japan: Japanese) and (1) SCZ in the Japanese population,⁶ (2) SCZ in multiple EAS populations (PGC Asia²⁵), and (3) a “combined” sample of (1) and (2) (current results). However, based on these comparisons, we did not detect any clear difference in r_g among these three subgroups (Table S2 and Figs. S3 and S4), specifically for the significant genetic correlation in six nonpsychiatric traits, except CRP (note: the genetic correlation of CRP was not significant in the EAS samples). This indicates that our comparison minimally affects the results.

Conclusion

Following LDSC analysis, the current study detected a correlation between MetS-related traits and the inflammatory marker CRP and SCZ in multipopulation samples. Although most of the statistically significant data were derived from EUR samples, we believe that novel and/or replicable relationships will be discovered in EAS samples by increasing the sample size. Therefore, further GWAS with larger sample sizes from multiple populations should provide unexpected relationships between psychiatric disorders and other phenotypes, and hopefully clarify the pathophysiology of psychiatric disorders.

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Author contributions

R.A., T.S., M.I., and N.I. contributed to the conception and study design. R.A., M.I., T.S., K.N., A.S., T.A., K.I., and N.I. provide substantial contributions to the analysis and interpretation of clinical data. R.A., M.I., and T.S. wrote the first draft of the article. All authors have contributed to and approved the final version of the article.

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Supporting information

Additional Supporting Information may be found in the online version of this article at the publisher's web-site:

Supplement Figure S1: Heat-map of genetic correlation analysis of binary traits for overall samples

Supplement Figure S2: Heat-map of genetic correlation analysis of quantitative trait loci (QTL) traits for overall samples

Supplement Figure S3: Heat-map of genetic correlation analysis of binary traits within East Asian samples

Supplement Figure S4: Heat-map of genetic correlation analysis of quantitative trait loci (QTL) traits within East Asian samples

Supplement Table S1: Results of genetic correlation analysis

Supplement Table S2: Results of genetic correlation analysis within East Asian samples